

RANDOM MINIMUM LENGTH SPANNING TREES IN REGULAR GRAPHS

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Consider a connected r -regular n -vertex graph G with random independent edge lengths, each uniformly distributed on $(0, 1)$. Let $mst(G)$ be the expected length of a minimum spanning tree. We show that $mst(G)$ can be estimated quite accurately under two distinct circumstances. Firstly, if r is large and G has a modest edge expansion property then $mst(G) \sim \frac{n}{r} \zeta(3)$, where $\zeta(3) = \sum_{j=1}^{\infty} j^{-3} \sim 1.202$. Secondly, if G has large girth then there exists an explicitly defined constant c_r such that $mst(G) \sim c_r n$. We find in particular that $c_3 = 9/2 - 6 \log 2 \sim 0.341$.

1. Introduction

Given a graph $G = (V, E)$ with edge lengths $\mathbf{x} = (x_e : e \in E)$, let $msf(G, \mathbf{x})$ denote the minimum length of a spanning forest. When $\mathbf{X} = (X_e : e \in E)$ is a family of independent random variables, each uniformly distributed on the interval $(0, 1)$, denote the expected value $\mathbf{E}(msf(G, \mathbf{X}))$ by $mst(G)$. This quantity gives a measure of the connectivity of G . In the most important case when G is connected, we use mst in place of msf in order to indicate *minimum spanning tree*.

Consider the complete graph K_n and the complete bipartite graph $K_{n,n}$. It is known (see [4, 5]) that, as $n \rightarrow \infty$, $mst(K_n) \rightarrow \zeta(3)$ and $mst(K_{n,n}) \rightarrow 2\zeta(3)$. Here $\zeta(3) = \sum_{j=1}^{\infty} j^{-3} \sim 1.202$. Also, it has recently been shown [12] that, for the d -cube Q_d , which has 2^d nodes and is regular of degree d , we have $(d/2^d)mst(Q_d) \rightarrow \zeta(3)$ as $d \rightarrow \infty$.

The results about mst quoted above (and others from [5]) are for particular regular graphs with growing degrees, and show that mst is about $\zeta(3)$ times the number of nodes divided by the degree. The results below provide a generalisation of all these results about mst . The first result gives a rather general lower bound. Let $\delta(G)$ and $\Delta(G)$ denote the minimum and maximum degree respectively of the graph G .

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Theorem 1. For any n -vertex graph G with no isolated vertices,

$$msf(G) \geq (1 + o(1))(n/\Delta)\zeta(3)$$

where $\Delta = \Delta(G)$ and the $o(1)$ term is with respect to $\Delta \rightarrow \infty$. In other words, for any $\varepsilon > 0$ there exist Δ_0 such that, for any graph G with no isolated vertices and with $\Delta = \Delta(G) \geq \Delta_0$, we have

$$msf(G) \geq (1 - \varepsilon)(n/\Delta)\zeta(3).$$

The above result in fact gives the right value for graphs $G = (V, E)$ that are regular or nearly regular and have a modest edge expansion property. For $S \subseteq V$, let $(S : \bar{S})$ be the set of edges with one end in S and the other in $\bar{S} = V \setminus S$.

Theorem 2. Let $\alpha = \alpha(r) = O(r^{-\frac{1}{3}})$ and let $\rho = \rho(r)$ and $\omega = \omega(r)$ tend to infinity with r . Suppose that the graph $G = (V, E)$ satisfies

$$(1) \quad r \leq \delta(G) \leq \Delta(G) \leq (1 + \alpha)r,$$

and

$$(2) \quad |(S : \bar{S})|/|S| \geq \omega r^{2/3} \log r \text{ for all } S \subseteq V \text{ with } r/2 < |S| \leq \min\{\rho r, |V|/2\}.$$

Then

$$msf(G) = (1 + o(1)) \frac{|V|}{r} \zeta(3)$$

where the $o(1)$ term is with respect to $r \rightarrow \infty$.

Note that for $|S| = k$ we have

$$(3) \quad |S : \bar{S}|/|S| \geq \delta - k + 1$$

and so we are really getting some expansion here for $|S| \leq \min\{\rho r, |V|/2\}$.

For regular graphs we of course take $\alpha = 0$. For K_n , $K_{n,n}$ and Q_d we can define ω, ρ such that the condition (2) holds: when $G = Q_d$ we use the result that

$$(4) \quad |(S : \bar{S})|/|S| \geq d - \log_2 |S|,$$

see for example Bollobás and Leader [3].

There are further similar results. Let $[d]$ denote the set $\{1, \dots, d\}$. Consider the d -dimensional mesh $M_{d,n}^{(1)} = (V_{d,n}, E_{d,n}^{(1)})$, where the vertex set $V_{d,n} = \{0, 1, \dots, n-1\}^d$ and if $x, y \in V_{d,n}$ then $\{x, y\}$ is in the edge set $E_{d,n}^{(1)}$ if and only if there exists $j \in [d]$ such that $x_i = y_i$ if $i \neq j$ and $x_j - y_j = \pm 1$. Thus $M_{d,n}^{(1)}$ has n^d vertices and has maximum degree $2d$ for $n \geq 3$. We also consider the ‘wrap-around’ version $M_{d,n}^{(2)} = (V_{d,n}, E_{d,n}^{(2)})$, where if $x, y \in V_{d,n}$ then $\{x, y\} \in E_{d,n}^{(2)}$ if and only if there exists $j \in [n]$ such that $x_i = y_i$ if $i \neq j$ and $x_j - y_j = \pm 1 \pmod n$. Thus $M_{d,n}^{(2)}$ is $2d$ -regular for $n \geq 3$. Both $M_{d,2}^{(1)}$ and $M_{d,2}^{(2)}$ are the d -cube Q_d , which is d -regular. The first part of the theorem below is Penrose’s result on the d -cube mentioned above.

Theorem 3. *If $d \rightarrow \infty$ then*

$$mst(Q_d) \sim \frac{2^d}{d} \zeta(3),$$

$$mst(M_{d,n}^{(2)}) \sim \frac{n^d}{2d} \zeta(3)$$

uniformly over $n \geq 3$, and if also $n \rightarrow \infty$ in such a way that $d = o(n)$ then

$$mst(M_{d,n}^{(1)}) \sim \frac{n^d}{2d} \zeta(3).$$

We now move on to discuss the second circumstance under which we can estimate $mst(G)$ quite accurately. Instead of considering graphs with large degrees, we consider r -regular graphs with large girth, or at least with few edges on short cycles. Recall that the girth of a graph G is the length of a shortest cycle in G .

Theorem 4. *For $r \geq 2$ let*

$$c_r = \frac{r}{(r-1)^2} \sum_{k=1}^{\infty} \frac{1}{k(k+\rho)(k+2\rho)},$$

where $\rho = 1/(r-1)$. Then, for any $r \geq 2$ and any r -regular graph G

$$|mst(G) - c_r n| \leq \frac{3n}{2g},$$

where n denotes the number of vertices and g denotes the girth of G . The constants c_r satisfy $c_2 = \frac{1}{2}$, $c_3 = 9/2 - 6\log 2 \sim 0.341$, $c_4 = 9 - 3\log 3 - \pi\sqrt{3} \sim 0.264$, and $c_5 = 15 - 10\log 2 - 5\pi/2 \sim 0.215$; and $c_r \sim \zeta(3)/r$ as $r \rightarrow \infty$.

Corollary 5. *For each $r \geq 2$ and $g \geq 3$, there exists $\delta = \delta(r, g) > 0$ with the following property. For every r -regular graph G with n vertices such that there is a set of at most δn edges which hit all cycles of length less than g , we have*

$$|mst(G) - c_r n| \leq \frac{2n}{g}.$$

From this corollary, we obtain easily a result about random regular graphs.

Let $G_{n,r}$ denote a random r -regular graph with vertex set $\{1, \dots, n\}$. Let the random variable $L_{n,r}$ be the minimum length of a spanning forest of the random regular graph $G_{n,r}$ when it has independent edge lengths each uniformly distributed on $(0, 1)$. Thus in the notation above we may write $L_{n,r} = mst(G_{n,r}, \mathbf{X})$ and $\mathbf{E}(L_{n,r}) = \mathbf{E}(mst(G_{n,r}))$.

Using the configuration model of random regular graphs see e.g. [2], it can easily be proved that

$$\Pr(G_{n,r} \text{ contains } \geq n^{1/2} \text{ edges on cycles of length } \leq \sqrt{\log n}) \leq n^{-(1/2-o(1))}.$$

We therefore have

Corollary 6. *For each integer $r \geq 3$,*

$$(1/n)\mathbf{E}(L_{n,r}) \rightarrow c_r.$$

Remark. Since for $r \geq 3$, $G_{n,r}$ is connected with probability $1 - O(n^{-2})$, this result is not changed if we condition on $G_{n,r}$ being connected.

Further information on the constants c_r is given in Propositions 10 and 11 below.

It is straightforward to extend these results to more general distributions on the edge lengths — see [5].

We also prove some results about how concentrated $mst(G, \mathbf{X})$ is about its mean.

Theorem 7.

(a) *For any r -regular graph $G = (V, E)$ with n vertices and $r = o((n/\log n)^{1/2})$,*

$$\Pr(|mst(G, \mathbf{X}) - mst(G)| \geq \varepsilon n/r) \leq e^{-\varepsilon^2 n/(5r^2)}$$

if n is sufficiently large.

(b) *There is a constant $K > 0$ such that the following holds. Suppose that*

$$|(S : \bar{S})| \geq \gamma r |S| \text{ for all } S \subset V \text{ with } |S| \leq n/2.$$

Then for any $0 < \varepsilon \leq 1$,

$$\Pr(|mst(G, \mathbf{X}) - mst(G)| \geq \varepsilon n/r) \leq n^2 e^{-K\varepsilon^2 \gamma^2 n/(\log n)^2},$$

for n sufficiently large.

The following two propositions are easier than Corollary 6, and have short proofs. The first concerns random 2-regular graphs, where we can give a more precise result than for general r .

Proposition 8.

$$\mathbf{E}(L_{n,2}) = n/2 - \log n + O(\sqrt{\log n}).$$

Finally, let us consider random graphs $G_{n,p}$ which are not too sparse. Consider any edge-probability $p = p(n)$ which is above the connectivity threshold, that is $P(G_{n,p} \text{ connected}) \rightarrow 1$ as $n \rightarrow \infty$. (Thus we are assuming that $p(n) = \frac{1}{2}n(\log n + \omega(n))$ where $\omega(n) \rightarrow \infty$ as $n \rightarrow \infty$.)

Proposition 9. *If $p = p(n)$ is above the threshold for connectivity, then $pmsf(G_{n,p}) \rightarrow \zeta(3)$ as $n \rightarrow \infty$, in probability and in any mean.*

2. Proofs

Given a graph $G=(V,E)$ with $|V|=n$ and $0\leq p\leq 1$, let G_p be the random subgraph of G with the same vertex set which contains those edges e with $X_e\leq p$. [Here we are assuming that as before we have a family $\mathbf{X}=(X_e:e\in E)$ of independent random variables each uniformly distributed on $(0,1)$.] Note that the edges of G are included independently with probability p . In this notation, the usual random graph $G_{n,p}$ could be written as $(K_n)_p$. Let $\kappa(G)$ denote the number of components of G . We shall first give a rather precise description of $msf(G)$.

Lemma 1. *For any graph G ,*

$$(5) \quad msf(G) = \int_{p=0}^1 \mathbf{E}(\kappa(G_p))dp - \kappa(G).$$

Proof. We shall follow the proof method in [1] and [7]. Let F denote the random set of edges in the minimal spanning forest. For any $0\leq p\leq 1$, $\sum_{e\in F} 1_{(X_e>p)}$ is the number of edges of F which are not in G_p , which equals $\kappa(G_p) - \kappa(G)$. But

$$msf(G, \mathbf{X}) = \sum_{e\in F} X_e = \sum_{e\in F} \int_{p=0}^1 1_{(X_e>p)} dp = \int_{p=0}^1 \sum_{e\in F} 1_{(X_e>p)} dp.$$

Hence

$$msf(G, \mathbf{X}) = \int_{p=0}^1 \kappa(G_p) dp - \kappa(G),$$

and the result follows on taking expectations. ■

2.1. Large Degrees

We substitute $p=x/r$ in (5) to obtain

$$msf(G) = \frac{1}{r} \int_{x=0}^r \mathbf{E}(\kappa(G_{x/r}))dx - \kappa(G).$$

Now let $C_{k,x}$ denote the total number of components in $G_{x/r}$ with k vertices. Thus

$$(6) \quad msf(G) = \frac{1}{r} \int_{x=0}^r \sum_{k=1}^n \mathbf{E}(C_{k,x})dx - \kappa(G).$$

We decompose

$$C_{k,x} = \tau_{k,x} + \sigma_{k,x}$$

where

$\tau_{k,x}$ denotes the number of tree components of $G_{x/r}$ with k vertices

and

$\sigma_{k,x}$ denotes the number of non-tree components in $G_{x/r}$ with k vertices.

We will find, perhaps not unexpectedly, that the number of components of $G_{x/r}$ is usually dominated by the number of components which are small trees. Imagine taking all trees T in G which have k vertices and giving them a root. Fix a vertex $v \in V$ and let $\mathcal{T}(v, k)$ be the set of trees obtained in this way which have root v . Let $t(v, k) = |\mathcal{T}(v, k)|$.

Lemma 2.

$$\frac{k^{k-2}(\delta - k)^{k-1}}{(k-1)!} \leq t(v, k) \leq \frac{k^{k-2}\Delta^{k-1}}{(k-1)!}.$$

Proof. Given a tree $T \in \mathcal{T}(v, k)$ we label v with k and then define a labelling $f: V(T) \setminus \{v\} \rightarrow \{1, \dots, k-1\}$ of the remaining vertices. Now consider pairs (T, f) where $T \in \mathcal{T}(v, k)$ and f is such a labelling. Clearly each rooted $T \in \mathcal{T}(v, k)$ is in $(k-1)!$ such pairs. Furthermore each such pair defines a unique spanning tree T' of K_k , where (i, j) is an edge of T' if and only if there is an edge $\{x, y\}$ of T such that $f(x) = i$ and $f(y) = j$. Each spanning tree T' of K_k nodes lies in between $(\delta - k)^{k-1}$ and Δ^{k-1} such pairs. Take a fixed breadth first search of T' starting at k and on reaching vertex ℓ for the first time, define $f^{-1}(\ell)$. There will always be between $\delta - k$ and Δ choices. Thus

$$(\delta - k)^{k-1}k^{k-2} \leq \#\text{pairs}(T, f) = t(v, k)(k-1)! \leq \Delta^{k-1}k^{k-2}$$

and the lemma follows. ■

Now consider a fixed sub-tree T of G containing k vertices. Suppose that the vertices of T induce $a(T)$ edges in G , and the sum of their degrees in G is $b(T)$. Then the probability $\pi(x, T)$ that it forms a component of $G_{x/r}$ satisfies

$$(7) \quad \pi(x, T) = \left(\frac{x}{r}\right)^{k-1} \left(1 - \frac{x}{r}\right)^{b(T) - a(T) - k + 1}.$$

Also

$$(8) \quad k - 1 \leq a(T) \leq \binom{k}{2} \quad \text{and} \quad k\delta \leq b(T) \leq k\Delta.$$

It follows from Lemma 2, (7) and (8) that

$$(9) \quad \mathbf{E}(\tau_{k,x}) \leq \frac{1}{k} \sum_v t(v, k) \left(\frac{x}{r}\right)^{k-1} \left(1 - \frac{x}{r}\right)^{k\delta - (k+2)(k-1)/2}$$

$$(10) \quad \leq \frac{nk^{k-2}}{k!} \left(\frac{\Delta}{r}\right)^{k-1} x^{k-1} \left(1 - \frac{x}{r}\right)^{k\delta - k^2}.$$

Similarly,

$$(11) \quad \mathbf{E}(\tau_{k,x}) \geq \frac{1}{k} \sum_v t(v, k) \left(\frac{x}{r}\right)^{k-1} \left(1 - \frac{x}{r}\right)^{k\Delta - 2k+2}$$

$$(12) \quad \geq \frac{nk^{k-2}}{k!} \left(\frac{\delta - k}{r}\right)^{k-1} x^{k-1} \left(1 - \frac{x}{r}\right)^{k\Delta}.$$

The $1/k$ factor in front of the sums in (9) and (11) comes from the fact that each k -vertex tree appears k times in the sum $\sum_v t(v, k)$. The following will be needed below:

$$\int_{x=0}^{\infty} x^{k-1} e^{-kx} dx = \frac{(k-1)!}{k^k} \geq \frac{1}{ke^k},$$

and for $a \geq 1$

$$\int_{x=a}^{\infty} x^{k-1} e^{-kx} dx \leq \int_{x=a}^{\infty} (xe^{-x})^k dx \leq \int_{x=a}^{\infty} e^{-kx/2} dx = \frac{2}{k} e^{-ka/2}.$$

Now, if $a, b \rightarrow \infty$, then

$$(13) \quad \int_{x=0}^a \sum_{k=1}^b \frac{k^{k-3}}{(k-1)!} x^{k-1} e^{-kx} dx = (1 + o(1)) \sum_{k=1}^b \frac{1}{k^3} \\ = (1 + o(1)) \zeta(3).$$

We may now prove Theorem 1: after that we shall continue the development here to prove Theorem 2.

Proof of Theorem 1. We use four stages.

(a) Let $\varepsilon > 0$. Let a and b be sufficiently large that

$$\int_{x=0}^a \sum_{k=1}^b \frac{k^{k-3}}{(k-1)!} x^{k-1} e^{-kx} dx \geq (1 - \varepsilon) \zeta(3).$$

Now, if $0 \leq x \leq r/2$ and $0 \leq \alpha \leq 1/2$, then

$$(1 - x/r)^{kr(1+\alpha)} \geq \exp\left(-k(1+\alpha)(x + 2x^2/r)\right) \geq e^{-kx} \exp(-xk\alpha - 3x^2k/r).$$

Let r_0 be sufficiently large that for $r \geq r_0$ we have $(1 - b/r)^{b-1} \geq (1 - \varepsilon)$ and $\exp(-3a^2b/r) \geq (1 - \varepsilon)$. Let $0 < \eta < 1/2$ be sufficiently small that $\exp(-ab\eta) \geq (1 - \varepsilon)$.

Now suppose that $r \geq r_0$, that the graph G has $\delta = \delta(G) = r$, and that $\Delta = \Delta(G) \leq (1 + \eta)r$. Then by (12) and the above, for $0 \leq x \leq a$ and $1 \leq k \leq b$,

$$\mathbf{E}(\tau_{k,x}) \geq \frac{n}{k} \frac{k^{k-2}}{(k-1)!} \left(1 - \frac{k}{r}\right)^{k-1} x^{k-1} \left(1 - \frac{x}{r}\right)^{k\Delta} \geq \frac{n}{k} \frac{k^{k-2}}{(k-1)!} x^{k-1} e^{-kx} (1 - \varepsilon)^3.$$

Hence $\text{msf}(G) \geq (1 - \varepsilon)^4 \frac{n}{r} \zeta(3)$.

(b) Next we drop the assumption on $\delta(G)$. Let $\varepsilon > 0$. We shall show that there exist r_1 and $\beta > 0$ such that, for any connected n -vertex graph G with $r_1 \leq \Delta = \Delta(G) \leq \beta n$, we have

$$\text{mst}(G) \geq (1 - \varepsilon) \frac{n}{\Delta} \zeta(3).$$

To do this, let r_0 and $\eta > 0$ be such that for any $r \geq r_0$ and any graph G with $\delta = \delta(G) = r$ and $\Delta = \Delta(G) \leq (1 + \eta)r$, we have $\text{msf}(G) \geq (1 - \varepsilon) \zeta(3)$. We have just seen that this is possible. Let $r_1 = \max\{r_0, 2/\eta\}$ and $\beta > 0$ be such that if $r_1 \leq r \leq \beta n$ then $r + \frac{r^2}{n-r} + 1 \leq (1 + \eta)r$.

Now let G be a connected n -vertex graph with $r_1 \leq r = \Delta(G) \leq \beta n$. We shall add edges to G to produce a graph G' which has minimum degree r and maximum degree $\Delta' \leq (1 + \eta)r$: then

$$\text{mst}(G) \geq \text{mst}(G') \geq (1 - \varepsilon) \frac{n}{\Delta'} \zeta(3),$$

and the desired result follows. To get G' we add edges between vertices of degree less than r until the vertices S of degree less than r form a clique. We then add new edges from S to $\bar{S} = V \setminus S$ until the vertices in S have degree r . When adding an $(S : \bar{S})$ edge we choose a vertex of current smallest degree in \bar{S} . In this way we end up with $\delta(G') = r$ and

$$\Delta' \leq r + \frac{r^2}{n-r} + 1 \leq (1 + \eta)r,$$

as required.

(c) Next we shall deduce the corresponding result for connected graphs but without the condition that $\Delta \leq \beta n$.

Let $\varepsilon > 0$. Choose r_1 and $\beta > 0$ as above for $\varepsilon/3$. Let r_2 be the maximum of r_1 and $\lceil 6/\varepsilon \rceil$. Consider a connected n -vertex graph G with $\Delta = \Delta(G) \geq r_2$. Let $k = \lceil (2/\beta) \rceil$, and form k disjoint copies G_1, \dots, G_k of G . For each $i = 1, \dots, k-1$ add a perfect matching between G_i and G_{i+1} . The new graph H is connected, and has

kn vertices and maximum degree $\Delta + 2$, and thus satisfies $\Delta(H) \leq 2n \leq \beta|V(H)|$. Hence

$$mst(H) \geq (1 - \varepsilon/3)(kn/(\Delta + 2))\zeta(3) \geq (1 - 2\varepsilon/3)(kn/\Delta)\zeta(3),$$

since $2/(\Delta + 2) < 2/r_1 \leq \varepsilon/3$. But $mst(H) \leq k mst(G) + (k - 1)/(n + 1)$, and so

$$mst(G) \geq (1/k)mst(H) - 1/n \geq (1 - 2\varepsilon/3)(n/\Delta)\zeta(3) - 1/n \geq (1 - \varepsilon)(n/\Delta)\zeta(3),$$

for $n \geq 3/\varepsilon$.

(d) Finally we remove the assumption of connectedness. Let c be the infimum of $mst(K_n)$ over all positive integers n . Then $c > 0$ — indeed it is easy to see that $c \geq 1/2$. Let $\varepsilon > 0$. Let r_2 be as above, and let r_3 be the maximum of r_2 and $\lceil \zeta(3)r_2/c \rceil$. Consider a graph G with $\Delta = \Delta(G) \geq r_3$. List the components of G as G_1, \dots, G_k where $G_i = (V_i, E_i)$. If $|V_i| < r_2$ then

$$mst(G_i) \geq c \geq r_2\zeta(3)/r_3 \geq |V_i|\zeta(3)/\Delta(G),$$

and if $|V_i| \geq r_2$ then

$$mst(G_i) \geq (1 - \varepsilon)|V_i|\zeta(3)/\Delta(G).$$

Hence

$$mst(G) = \sum_{i=1}^k mst(G_i) \geq (1 - \varepsilon) \left(\sum_{i=1}^k |V_i| \right) \zeta(3) / \Delta(G) = (1 - \varepsilon) |V(G)| \zeta(3) / \Delta(G),$$

as required. This completes the proof of Theorem 1. ■

Proof of Theorem 2.

In order to use (6) we need to consider a number of separate ranges for x and k . Let $A = 2r^{1/3}/\omega$, $B = \lfloor (Ar)^{1/4} \rfloor$ so that each of $B\alpha$, AB^2/r and $A/B \rightarrow 0$ as $r \rightarrow \infty$.

Range 1. $0 \leq x \leq A$ and $1 \leq k \leq B$. By (10) we have

$$\mathbf{E}(\tau_{k,x}) \leq \frac{nk^{k-2}}{k!} x^{k-1} e^{-kx} \exp(k\alpha + xk^2/r),$$

since $(\Delta/r)^{k-1} \leq (1 + \alpha)^k \leq \exp(k\alpha)$, and $(1 - x/r)^{k\delta - k^2} \leq \exp(-xk + xk^2/r)$. Also,

$$\exp(k\alpha + xk^2/r) \leq \exp(B\alpha + AB^2/r) = 1 + o(1).$$

Hence

$$\begin{aligned} \frac{1}{r} \int_{x=0}^A \sum_{k=1}^B \mathbf{E}(\tau_{k,x}) dx &\leq (1 + o(1)) \frac{n}{r} \int_{x=0}^A \sum_{k=1}^B \frac{k^{k-2}}{k!} x^{k-1} e^{-kx} dx \\ (14) \qquad \qquad \qquad &\leq (1 + o(1)) \frac{n}{r} \zeta(3). \end{aligned}$$

Let $\sigma_{k,u,x}$ be the number of non-tree components of $G_{x/r}$ which have k vertices and $k-1+u$ edges. Then

$$\mathbf{E}(\sigma_{k,u,x}) \leq \frac{1}{k} \sum_{v \in V} t(v, k) \binom{k}{2}^u \left(\frac{x}{r}\right)^{k-1+u} \left(1 - \frac{x}{r}\right)^{kr-k^2}.$$

So

$$\begin{aligned} \mathbf{E}(\sigma_{k,x}) &\leq \frac{nk^{k-2}}{k!} \Delta^{k-1} \sum_{u=1}^{\infty} \left(\frac{k^2}{2}\right)^u \left(\frac{x}{r}\right)^{k-1+u} \left(1 - \frac{x}{r}\right)^{kr-k^2} \\ &\leq \frac{nk^{k-2}}{k!} \left(\frac{\Delta}{r}\right)^{k-1} x^{k-1} e^{-xk} e^{xk^2/r} \sum_{u=1}^{\infty} \left(\frac{k^2 x}{2r}\right)^u \\ &\leq \left(\frac{e^{k\alpha+xk^2/r}}{2-xk^2/r}\right) \frac{n}{r} \frac{k^k}{k!} x^k e^{-kx} \\ &\leq \frac{n}{r} \frac{k^k}{k!} x^k e^{-kx} \end{aligned}$$

if r is sufficiently large. Thus,

$$\begin{aligned} \frac{1}{r} \int_{x=0}^A \sum_{k=1}^B \mathbf{E}(\sigma_{k,x}) &\leq \frac{n}{r^2} \sum_{k=1}^B \frac{k^k}{k!} \int_{x=0}^{\infty} x^k e^{-kx} dx \\ &= \frac{n}{r^2} \sum_{k=1}^B \frac{1}{k^2} \\ (15) \qquad \qquad \qquad &\leq 2 \frac{n}{r^2} = o(n/r). \end{aligned}$$

Range 2. $x \leq A$ and $k \geq B$. Using the bound

$$(16) \qquad \qquad \qquad \sum_{k=\ell}^n C_{k,x} \leq \frac{n}{\ell}$$

for all ℓ, x we get

$$(17) \qquad \frac{1}{r} \int_{x=0}^A \sum_{k=B}^n \mathbf{E}(C_{k,x}) dx \leq \frac{1}{r} \int_{x=0}^A \frac{n}{B} dx = \frac{A}{B} \cdot \frac{n}{r} = o(n/r).$$

We next have to consider larger values of x in our integral. Now G contains at most $n(e\Delta)^k$ connected subgraphs with k vertices. To see this, choose $v \in V$

and note that G contains fewer than $(e\Delta)^k$ k -vertex trees rooted at v . This follows from the formula (29) below for the number of subtrees of an infinite rooted r -ary tree which contain the root.

Also, from (3) we get $S \subseteq V$, $|S|=k$ implies $|S:\bar{S}| \geq k\delta - k(k-1) \geq k(r-k)$. Thus

$$\begin{aligned} \mathbf{E}(C_{k,x}) &\leq n(e\Delta)^k \left(1 - \frac{x}{r}\right)^{k(r-k)} \\ (18) \qquad &\leq n(re^{1+\alpha-x(1-k/r)})^k. \end{aligned}$$

Range 3. $x \geq A$ and $k \leq r/2$. Equation (18) implies that for large r ,

$$(19) \qquad \mathbf{E}(C_{k,x}) \leq ne^{-kA/3}.$$

Thus

$$(20) \qquad \frac{1}{r} \int_A^r \sum_{k=1}^{r/2} \mathbf{E}(C_{k,x}) dx \leq nre^{-A/3} = o(n/r).$$

Range 4. $x \geq A$ and $r/2 < k \leq k_0 = \min\{\rho r, n/2\}$. It is only here that we use the expansion condition (2). We find

$$(21) \qquad \mathbf{E}(C_{k,x}) \leq n(er)^k \left(1 - \frac{x}{r}\right)^{k\omega r^{2/3} \log r} \leq n \left(\frac{e}{r}\right)^k.$$

So,

$$(22) \qquad \frac{1}{r} \int_A^r \sum_{k=r/2+1}^{k_0} \mathbf{E}(C_{k,x}) dx \leq n \left(\frac{e}{r}\right)^{r/2} = o(n/r).$$

We split the remaining range into two cases.

Range 5. $x \geq A$ and $k > k_0$.

Case 1. $n \geq 2\rho r$, so that $k_0 = \rho r$.

If $k \geq k_0$ we use (16) to deduce that

$$(23) \qquad \frac{1}{r} \int_A^r \sum_{k=\rho r}^n \mathbf{E}(C_{k,x}) dx \leq \frac{n}{\rho r} = o(n/r).$$

Part (b) now follows from (6), (14), (15), (17), (20), (22) and (23).

Case 2. $n < 2\rho r$, so that $k_0 = n/2$.

For larger r , we have to use the $-\kappa(G)$ term in (6), ignored in the previous case. Here (2) implies $\kappa(G)=1$. We deduce from (19) and (21) that

(24)
$$\Pr(G_{A/r} \text{ is not connected}) \leq 2ne^{-A/3} + 2n \left(\frac{e}{r}\right)^{r/2}.$$

Then,

$$\frac{1}{r} \int_{x=A}^r \sum_{k=0}^n \mathbf{E}(C_{k,x})dx = 1 - O(n^{-K})$$

for any constant $K>0$, and the proof is completed by (6), (14), (15), (17). ■

Remark. It is worth pointing out that it is not enough to have $r \rightarrow \infty$ in order to have Theorem 2, that is, we need some extra condition such as the expansion condition (2). For consider the graph G_0 obtained from n/r r -cliques $C_1, C_2, \dots, C_{n/r}$ by deleting an edge (x_i, y_i) from C_i , $1 \leq i \leq n/r$ then joining the cliques into a cycle of cliques by adding edges (y_i, x_{i+1}) for $1 \leq i \leq n/r$. It is not hard to see that

$$mst(G_0) \sim \frac{n}{r} \left(\zeta(3) + \frac{1}{2} \right)$$

if $r \rightarrow \infty$ with $r=o(n)$. We conjecture that this is the worst-case, that is

Conjecture. *Assuming only the conditions of Theorem 1,*

$$mst(G) \leq (1 + o(1)) \frac{n}{r} \left(\zeta(3) + \frac{1}{2} \right).$$

2.1.1. Proof of Theorem 3

We consider $M_{d,n}^{(2)}$ first. We prove the equivalent of (4). For this we need a technical lemma.

Lemma 3. *Assume $s_1, s_2, \dots, s_n \geq 0$ and $s = s_1 + s_2 + \dots + s_n$ then*

(25)
$$\frac{1}{2} s \log_2 s \geq \frac{1}{2} \sum_{i=1}^n s_i \log_2 s_i + \sum_{i=1}^n \min\{s_i, s_{i+1}\}.$$

(Here $s_{n+1} = s_1$ and $s_i \log_2 s_i = 0$ when $s_i = 0$.)

Proof. We prove (25) by induction on n . The case $n = 2$ is proved in [3]. Assume (25) is true for some $n \geq 2$ and consider $n + 1$.

$$\begin{aligned} L &= \frac{1}{2} \sum_{i=1}^{n+1} s_i \log_2 s_i + \sum_{i=1}^{n+1} \min\{s_i, s_{i+1}\} \\ &\leq \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} \\ &\quad + \min\{s_n, s_{n+1}\} + \min\{s_{n+1}, s_1\} - \min\{s_n, s_1\} \end{aligned}$$

by induction.

Case 1. $\min\{s_1, s_n, s_{n+1}\} = s_1$:

$$\begin{aligned} L &\leq \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} + \min\{s_n, s_{n+1}\} \\ &\leq \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} + \min\{s - s_{n+1}, s_{n+1}\} \\ &\leq \frac{1}{2}s \log_2 s. \end{aligned}$$

Case 2. $\min\{s_1, s_n, s_{n+1}\} = s_n$: similar.

Case 3. $\min\{s_1, s_n, s_{n+1}\} = s_{n+1}$:

$$\begin{aligned} L &\leq \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} + 2s_{n+1} - \min\{s_n, s_1\} \\ &\leq \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} + s_{n+1} \\ &= \frac{1}{2}(s - s_{n+1}) \log_2(s - s_{n+1}) + \frac{1}{2}s_{n+1} \log_2 s_{n+1} + \min\{s - s_{n+1}, s_{n+1}\} \\ &\leq \frac{1}{2}s \log_2 s. \end{aligned} \quad \blacksquare$$

Now consider $S \subseteq V_{d,n}$ with $|S| = s$. We now prove by induction on s that

$$(26) \quad S \text{ contains at most } \frac{1}{2}s \log_2 s \text{ edges.}$$

Let S_i be the set of vertices $x \in S$ with $x_n = i$. Let $s_i = |S_i|$, $i = 1, 2, \dots, n$. Each S_i can be considered a subset of $V_{d,n-1}$ and we can assume inductively that each S_i contains at most $\frac{1}{2}s_i \log_2 s_i$ edges. Therefore S contains at most L edges and (26) follows from Lemma 3. It follows that $|S : \bar{S}| \geq 2ds - s \log_2 s$ and so $M_{d,n}^{(2)}$ has adequate expansion to apply Theorem 2.

Now consider the spanning subgraph $M_{d,n}^{(1)}$ of $M_{d,n}^{(2)}$. Since each edge of $M_{d,n}^{(2)}$ is equally likely to be in a minimum spanning tree T , the expected number of ‘wrap-around’ edges in T equals $(n^d - 1)/n < n^{d-1}$. Hence

$$mst(M_{d,n}^{(2)}) \leq mst(M_{d,n}^{(1)}) \leq mst(M_{d,n}^{(2)}) + n^{d-1},$$

which completes the proof. ■

2.2. Large Girth

We note first that all components of G_p with fewer than g vertices are trees. Here g denotes the girth of G . Hence

$$(27) \quad \left| mst(G) - \int_{p=0}^1 \sum_{k=1}^{g-1} \mathbf{E}(\tau_{k,p}) dp \right| \leq \frac{n}{g}.$$

Here $\tau_{k,p}$ is the number of (tree) components with k vertices in G_p and n/g is an upper bound for the number of components of G_p with g or more vertices.

Let $t(v, k)$ be as in Lemma 2. This time we have an exact formula for $t(v, k)$ when k is less than the girth g of G .

Lemma 4. For $k < g$,

$$t(v, k) = \frac{r((r-1)k)!}{(k-1)!((r-2)k+2)!}.$$

Proof. We use the formula

$$(28) \quad t(v, k) = \sum_{i=1}^k \frac{1}{(r-2)i+1} \binom{(r-1)i}{i} \frac{1}{(r-2)(k-i)+1} \binom{(r-1)(k-i)}{k-i}.$$

This follows from the formula

$$(29) \quad \frac{1}{(r-1)m+1} \binom{rm}{m}$$

for the number of m -vertex subtrees of an infinite rooted r -ary tree which contain the root — see Knuth [8], Problem 2.3.4.4.11. To obtain (28) we take each tree with k vertices rooted at v and view it as an $(r-1)$ -ary tree with i vertices rooted at v plus an $(r-1)$ -ary tree with $k-i$ vertices rooted at the largest (numbered) neighbour of v . Let

$$a_k = \sum_{i=0}^k \frac{1}{(r-2)i+1} \binom{(r-1)i}{i} \frac{1}{(r-2)(k-i)+1} \binom{(r-1)(k-i)}{k-i}.$$

[Sum from $i=0$ as opposed to $i=1$ in (28).] Then

$$\begin{aligned}
 & \sum_{k=0}^{\infty} a_k x^k \\
 &= \sum_{k=0}^{\infty} \sum_{i=0}^k \frac{1}{(r-2)i+1} \binom{(r-1)i}{i} \frac{1}{(r-2)(k-i)+1} \binom{(r-1)(k-i)}{k-i} x^k \\
 &= \sum_{i=0}^{\infty} \frac{1}{(r-2)i+1} \binom{(r-1)i}{i} x^i \sum_{k=i}^{\infty} \frac{1}{(r-2)(k-i)+1} \binom{(r-1)(k-i)}{k-i} x^{k-i} \\
 &= \left(\sum_{i=0}^{\infty} \frac{1}{(r-2)i+1} \binom{(r-1)i}{i} x^i \right)^2 \\
 &= \left(\sum_{i=0}^{\infty} \frac{1}{(r-1)i+1} \binom{(r-1)i+1}{i} x^i \right)^2 \\
 &= B_{r-1}(x)^2,
 \end{aligned}$$

where

$$B_t(x) = \sum_{i=0}^{\infty} \frac{1}{ti+1} \binom{ti+1}{i} x^i$$

is the *Generalised Binomial Series*. The identity

$$B_t(x)^s = \sum_{i=0}^{\infty} \frac{s}{ti+s} \binom{ti+s}{i} x^i$$

is given for example in Graham, Knuth and Patashnik [6]. Thus,

$$a_k = \frac{2}{(r-1)k+2} \binom{(r-1)k+2}{k}.$$

The lemma follows from

$$t(v, k) = a_k - \frac{1}{(r-2)k+1} \binom{(r-1)k}{k}.$$

■

We may now prove the first part of Theorem 4. We have

$$(30) \quad \int_{p=0}^1 \sum_{k=1}^{g-1} \mathbf{E}(\tau_{k,p}) dp$$

$$\begin{aligned}
 &= \frac{1}{k} \int_{p=0}^1 \sum_{k=1}^{g-1} \sum_{v \in V} t(v, k) p^{k-1} (1-p)^{rk-2k+2} dp \\
 &= \sum_{k=1}^{g-1} \frac{n}{k} \frac{r((r-1)k)!}{(k-1)!((r-2)k+2)!} \frac{(k-1)!((r-2)k+2)!}{((r-1)k+2)!} \\
 &= \sum_{k=1}^{g-1} \frac{nr}{k((r-1)k+1)((r-1)k+2)} \\
 &= \frac{nr}{(r-1)^2} \sum_{k=1}^{g-1} \frac{1}{k(k+\rho)(k+2\rho)}
 \end{aligned}$$

where $\rho = 1/(r-1)$. Theorem 4 now follows from (27) and

$$\begin{aligned}
 \frac{r}{(r-1)^2} \sum_{k=g}^{\infty} \frac{1}{k(k+\rho)(k+2\rho)} &\leq \frac{r}{(r-1)^2} \sum_{k=g}^{\infty} k^{-3} \\
 &\leq \frac{r}{(r-1)^2} \int_{g-1}^{\infty} x^{-3} dx \\
 &= \frac{r}{(r-1)^2} \frac{1}{2(g-1)^2} \\
 &\leq \frac{1}{2g}.
 \end{aligned}$$

Proof of Corollary 5. Start with a 2-edge-connected r -regular graph with girth at least $g-2$, and form a new graph H by ‘splitting’ an edge so that two vertices have degree 1 and all the others have degree r .

Let F be a set of edges in G which meet each cycle of length less than g . From the graph G , form a new graph \hat{G} as follows. For each edge $f = \{u, v\} \in F$, take a new copy H_f of H and identify the vertices u and v with the vertices of degree 1 in H_f . Then \hat{G} has girth at least g , $|V(\hat{G})| = n + |F|(|V(H)| - 2) = (1 + o(1))n$, and $|msf(\hat{G}) - msf(G)| \leq |F||E(H)| = o(n)$. ■

2.2.1. Proof of Theorem 7

Our main tool here is a concentration inequality of Talagrand [14], see Steele [13] for a good exposition. Let A be a (measurable) non-empty subset of \mathbb{R}^E . For

$\mathbf{x}, \beta \in \mathbb{R}^E$ with $\|\beta\|_2 = 1$ let

$$(31) \quad d_A(\mathbf{x}, \beta) = \inf_{\mathbf{y} \in A} \sum_{e \in E} \beta_e 1_{\{\mathbf{x}_e \neq \mathbf{y}_e\}}.$$

and let

$$d_A(\mathbf{x}) = \sup_{\beta} d_A(\mathbf{x}, \beta).$$

Talagrand shows that for all $t > 0$,

$$(32) \quad \Pr(\mathbf{X} \in A) \Pr(d_A(\mathbf{X}) \geq t) \leq e^{-t^2/4}.$$

(a) For $a \in \mathbb{R}$ let

$$S(a) = \{\mathbf{y} \in \mathbb{R}^E : mst(G, \mathbf{y}) \leq a\}.$$

Given \mathbf{x} we let $T = T(\mathbf{x})$ be a minimum spanning tree of G using these weights ($T(\mathbf{X})$ is unique with probability 1). Let $L = L(\mathbf{x}) = (\sum_{e \in T} \mathbf{x}_e^2)^{1/2}$. Note that $L(\mathbf{x}) \leq n^{1/2}$. Define, $\beta = \beta(\mathbf{x})$ by

$$\beta_e = \begin{cases} \mathbf{x}_e / L & : e \in T \\ 0 & : \text{otherwise.} \end{cases}$$

Then for $\mathbf{y} \in S(a)$ we have

$$\begin{aligned} mst(G, \mathbf{x}) &\leq mst(G, \mathbf{y}) + \sum_{e \in T(\mathbf{x})} (\mathbf{x}_e - \mathbf{y}_e)^+ \\ &\leq mst(G, \mathbf{y}) + L(\mathbf{x}) \sum_{e \in E} \beta_e 1_{\{\mathbf{x}_e \neq \mathbf{y}_e\}}. \end{aligned}$$

By choosing \mathbf{y} achieving the minimum in (31) (the infimum is achieved) we see that

$$mst(G, \mathbf{x}) \leq a + L(\mathbf{x}) d_a(\mathbf{x}, \beta) \leq a + n^{1/2} d_a(\mathbf{x}, \beta).$$

Applying (32) with $A = S(a)$ we get

$$(33) \quad \Pr(mst(G, \mathbf{X}) \leq a) \Pr(mst(G, \mathbf{X}) \geq a + n^{1/2} t) \leq e^{-t^2/4}.$$

Let M denote the median of $mst(G, \mathbf{X})$. Then with $a = M$ and $t = \varepsilon n^{1/2}/r$,

$$(34) \quad \Pr(mst(G, \mathbf{X}) \geq M + \varepsilon n/r) \leq 2e^{-\varepsilon^2 n/(4r^2)}.$$

With $a = M - \varepsilon n/r$,

$$(35) \quad \Pr(mst(G, \mathbf{X}) \leq M - \varepsilon n/r) \leq 2e^{-\varepsilon^2 n/(4r^2)}.$$

Equations (34) and (35) plus $r = o((n/\log n)^{1/2})$ imply that

$$|M - mst(G)| = o(n/r)$$

and so it is a simple matter to replace M by $mst(G)$ in (34), (35) to obtain (a).

(b) We change the definition of β slightly. For minimum spanning tree $T(\mathbf{x})$ we let $T_1(\mathbf{x}) = \{e \in T : \mathbf{x}_e \leq 12 \log n / (\gamma r)\}$. Then let

$$L_1(\mathbf{x}) = \left(\sum_{e \in T_1} \mathbf{x}_e^2 \right)^{1/2} \leq \frac{12n^{1/2} \log n}{\gamma r}.$$

Then define

$$\beta_e = \begin{cases} \mathbf{x}_e / L_1 & e \in T_1 \\ 0 & \text{otherwise} \end{cases}.$$

Also let

$$\phi(\mathbf{x}) = \sum_{e \in T \setminus T_1} \mathbf{x}_e.$$

Then for $\mathbf{y} \in S(a)$ we have

$$\begin{aligned} mst(G, \mathbf{x}) &\leq mst(G, \mathbf{y}) + \sum_{e \in T_1} (\mathbf{x}_e - \mathbf{y}_e)^+ + \phi(\mathbf{x}) \\ &\leq mst(G, \mathbf{y}) + L_1(\mathbf{x}) \sum_{e \in E} \beta_e \mathbf{1}_{\{\mathbf{x}_e \neq \mathbf{y}_e\}} + \phi(\mathbf{x}). \end{aligned}$$

By choosing \mathbf{y} achieving the minimum in (31) we see that

$$mst(G, \mathbf{x}) \leq a + L_1(\mathbf{x}) d_a(\mathbf{x}, \beta) + \phi(\mathbf{x}).$$

Applying (32) we get

$$(36) \quad \Pr(mst(G, \mathbf{X}) \leq a) \Pr(mst(G, \mathbf{X}) \geq a + t \frac{12n^{1/2} \log n}{\gamma r} + \phi(\mathbf{X})) \leq e^{-t^2/4}.$$

We will show below that

$$(37) \quad \Pr(\phi(\mathbf{X}) \geq \varepsilon n / (3r)) \leq e^{-\gamma n / (20(\log n)^2)}.$$

So putting $a = M$ and $t = \varepsilon \gamma n^{1/2} / (36 \log n)$ into (36) we get

$$\Pr(mst(G, \mathbf{X}) \geq M + 2\varepsilon n / (3r)) \leq 2e^{-\varepsilon^2 \gamma^2 n / (5184(\log n)^2)} + \Pr(\phi(\mathbf{X}) \geq \varepsilon n / (3r)).$$

On the other hand, putting $a = M - 2\varepsilon n/(3r)$ and $t = \varepsilon \gamma n^{1/2}/(36 \log n)$ we get

$$\Pr(\text{mst}(G, \mathbf{X}) \leq M - 2\varepsilon n/(3r)) \Pr(\text{mst}(G, \mathbf{X}) \geq M - \varepsilon n/(3r) + \phi(\mathbf{X})) \leq e^{-t^2/4}.$$

But

$$\Pr(\text{mst}(G, \mathbf{X}) \geq M - \varepsilon n/(3r) + \phi(\mathbf{X})) \geq \frac{1}{2} - \Pr(\phi(\mathbf{X}) \geq \varepsilon n/(3r))$$

and we can finish as in (a).

Proof of (37). Let

$$\begin{aligned} \pi(m, k, p) &= \Pr(G_p \text{ contains } \geq m \text{ components of size } k) \\ &\leq \binom{n}{k}^m (1-p)^{\gamma k r m/2} \\ &\leq \left(\frac{ne}{k} e^{-p\gamma r/2} \right)^{mk} \\ &\leq e^{-mkp\gamma r/3} \end{aligned}$$

if $p \geq p_0 = \min\{1, 12 \log n/(\gamma r)\}$. Next let

$$p_i = \min\{1, 2^i p_0\} \text{ for } 0 \leq i \leq i_0 = \lceil \log_2 p_0^{-1} \rceil$$

and

$$m_{k,p} = \frac{\varepsilon n}{6kpr(\log n)^2}.$$

Now

$$\phi(\mathbf{X}) \leq \sum_{i=0}^{i_0-1} \sum_{k=1}^n C_{k,p_i} p_{i+1}$$

and so if

$$(38) \quad G_{p_i} \text{ contains } < m_{k,p_i} \text{ components of size } k \text{ for } 0 \leq i < i_0, 1 \leq k \leq n$$

then

$$\phi(\mathbf{X}) \leq \sum_{i=0}^{i_0-1} \sum_{k=1}^n \frac{\varepsilon n}{3kr(\log n)^2} \leq \frac{\varepsilon n}{3r}.$$

Furthermore, the probability that (38) fails to hold is at most

$$\sum_{i=0}^{i_0-1} \sum_{k=1}^n \pi(m_{k,p_i}, k, p_i) \leq \sum_{i=0}^{i_0-1} \sum_{k=1}^n e^{-\varepsilon \gamma n/(18(\log n)^2)}$$

which proves (37). ■

We now consider the values of the constants c_r more carefully.

Proposition 5. *The constants c_r satisfy $c_2 = 1/2$, $c_3 = 9/2 - 6\log 2 \sim 0.341$, $c_4 = 9 - 3\log 3 - \pi\sqrt{3} \sim 0.264$ and $c_5 = 15 - 10\log 2 - 5\pi/2 \sim 0.215$; and in general, for $r \geq 3$, $c_r = r \sum_{j=0}^{r-2} g(\omega^j)$, where $g(x) = \frac{(x-1)^2}{2x^2} \log\left(\frac{1}{1-x}\right) + \frac{3}{4}$ and $\omega = e^{2\pi i/(r-1)}$.*

Proof. Let Σ_r denote the sum in Theorem 4, so that $c_r = r\Sigma_r$. Note first that

$$\begin{aligned} \sum_{k=1}^{\infty} \frac{x^k}{k(k+1)(k+2)} \\ &= \sum_{k=1}^{\infty} x^k \left(\frac{1}{2k} - \frac{1}{k+1} + \frac{1}{2(k+2)} \right) \\ &= \frac{1}{2} \log \left(\frac{1}{1-x} \right) - \frac{1}{x} \left(\log \left(\frac{1}{1-x} \right) - x \right) + \frac{1}{2x^2} \left(\log \left(\frac{1}{1-x} \right) - x - \frac{x^2}{2} \right) \\ &= \frac{(x-1)^2}{2x^2} \log \left(\frac{1}{1-x} \right) + \frac{3}{4} - \frac{1}{2x}. \end{aligned}$$

Thus $\Sigma_2 = \frac{1}{4}$. Also, for $r \geq 3$, note that $\omega^{r-1} = 1$ and $1 + \omega + \dots + \omega^{r-2} = 0$. Hence, for $r \geq 3$

$$\Sigma_r = (r-1) \sum_{k:(r-1) \nmid k} \frac{1}{k(k+1)(k+2)} = \sum_{j=0}^{r-2} g(\omega^j).$$

For $r=3$, $\omega = -1$ so

$$\Sigma_3 = g(1) + g(-1) = \frac{3}{2} - 2\log 2,$$

and thus c_3 is as given. For $r=4$ we find after some calculation that

$$Re(g(\omega)) = \frac{3}{4} - \frac{3\log 3}{8} - \frac{\pi\sqrt{3}}{8}.$$

But $\Sigma_4 = \frac{3}{4} + 2Re(g(\omega))$ and so c_4 is as given. For $r=5$, $\omega = i$ and we find that

$$\Sigma_5 = g(1) + g(i) + g(-1) + g(-i) = 3 - 2\log 2 - \pi/2,$$

and so c_5 is as given. ■

Proposition 6. *For any $r \geq 2$,*

$$\frac{\zeta(3)}{r+1} < c_r < \frac{r\zeta(3)}{(r-1)^2}.$$

Also

$$\begin{aligned} c_r &= r \sum_{k=3}^{\infty} \left(-\frac{1}{r-1} \right)^{k-1} (2^{k-2} - 1) \zeta(k) \\ &= \frac{r}{(r-1)^2} \zeta(3) - 3 \frac{r}{(r-1)^3} \zeta(4) + 7 \frac{r}{(r-1)^4} \zeta(5) - \dots \end{aligned}$$

Both of these results show that $c_r \sim \zeta(3)/r$ as $r \rightarrow \infty$.

Proof. We may write

$$c_r = r(r-1)^{-2} \sum_{k=1}^{\infty} (k(k+1/(r-1))(k+2/(r-1)))^{-1}.$$

It follows that $c_r < \frac{r}{(r-1)^2} \zeta(3)$, and

$$c_r > r(r-1)^{-2} \left(1 + \frac{1}{r-1}\right)^{-1} \left(1 + \frac{2}{r-1}\right)^{-1} \zeta(3) = \frac{1}{r+1} \zeta(3).$$

Also, for any $0 \leq x \leq 1$

$$\sum_{k=1}^{\infty} (k(k+x))^{-1} = \sum_{k=1}^{\infty} k^{-2} \sum_{j=0}^{\infty} (-x/k)^j = \sum_{k=2}^{\infty} (-x)^{k-2} \zeta(k).$$

Hence, for any $a > 1$

$$\begin{aligned} \sum_{k=1}^{\infty} (k(k+1/a)(k+2/a))^{-1} &= \sum_{k=1}^{\infty} \frac{a}{k} \left(\frac{1}{k+1/a} - \frac{1}{k+2/a} \right) \\ &= \sum_{k=3}^{\infty} (-1/a)^{k-3} (2^{k-2} - 1) \zeta(k). \end{aligned}$$

Thus

$$\begin{aligned} c_r &= r \sum_{k=3}^{\infty} \left(-\frac{1}{r-1} \right)^{k-1} (2^{k-2} - 1) \zeta(k) \\ &= \frac{r}{(r-1)^2} \zeta(3) - 3 \frac{r}{(r-1)^3} \zeta(4) + 7 \frac{r}{(r-1)^4} \zeta(5) - \dots \end{aligned}$$

It remains only to prove Propositions 8 and 9.

Proof of Proposition 8. We estimate the maximum total weight of edges that can be deleted without increasing the number of components, which are all cycles. Let C_k be the random number of k -cycles in $G_{n,2}$. Then using the configuration model we can prove that for $k \geq 3$, $\mathbf{E}(C_k) = (1 + O(\frac{k}{n})) \frac{1}{k}$. So the expected ‘savings’ from k -cycles is

$$\left(1 + O\left(\frac{k}{n}\right)\right) \frac{1}{k} \left(1 - \frac{1}{k+1}\right) = \left(1 + O\left(\frac{k}{n}\right)\right) \frac{1}{k+1}.$$

Hence the total expected savings from cycles of length at most k is

$$\left(1 + O\left(\frac{k}{n}\right)\right) \sum_{i=3}^k \frac{1}{i+1} = \left(1 + O\left(\frac{k}{n}\right)\right) (\log k + O(1)).$$

Take $k \sim n/\sqrt{\log n}$. Then the total savings is at least

$$\left(1 + O\left(\frac{k}{n}\right)\right) (\log k + O(1)) = \log n + O(\sqrt{\log n}),$$

and is at most this value plus $n/k \sim \sqrt{\log n}$. ■

Proof of Proposition 9. Consider the complete graph K_n , with independent edge lengths X_e on the edges e , each uniformly distributed on $(0,1)$. Call this the random network (K_n, \mathbf{X}) . Form a random subgraph H on the same set of vertices by including the edge e exactly when $X_e \leq p$, and give e the length X_e/p . We thus obtain a random graph $G_{n,p}$ with independent edge lengths, each uniformly distributed on $(0,1)$. Call this the random network (H, \mathbf{Y}) . We observe

$$mst(K_n, \mathbf{X}) \mathbf{1}_{(H \text{ connected})} \leq pmsf(H, \mathbf{Y}) \leq mst(K_n, \mathbf{X}).$$

The theorem now follows easily from the fact that that $mst(K_n, \mathbf{X}) \rightarrow \zeta(3)$ as $n \rightarrow \infty$, in probability and in any mean [4, 5]. ■

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